


RESEARCH ARTICLE

Mammographic density and exposure to air pollutants in premenopausal women: a cross-sectional study

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Abstract

Background: Mammographic density (MD) is a well-established risk factor for breast cancer. Air pollution is a major public health concern and a recognized carcinogen. We aim to investigate the association between MD and exposure to specific air pollutants (SO₂, CO, NO, NO₂, NO_x, PM_{2.5}, PM₁₀, and O₃) in premenopausal females.

Methods: This cross-sectional study, carried out in Spain, included 769 participants who attended their gynecological examinations. Hourly concentrations of the pollutants were extracted from the Air Quality Monitoring System of Madrid City over a 3-year period. Individual long-term exposure to pollutants was assessed by geocoding residential addresses and monitoring stations, and applying ordinary kriging to the 3-year annual mean concentrations of each pollutant to interpolate the surface of Madrid. This exposure variable was categorized into quartiles. In a first analysis, we used multiple linear regression models with the log-transformed percent MD as a continuous variable. In a second analysis, we used MD as a dichotomous variable (“high” density (MD > 50%) vs. “low” density (MD ≤ 50%)) and applied multiple logistic regression models to estimate odds ratios (ORs). We also analyzed the correlation among the pollutants, and performed a principal component analysis (PCA) to reduce the dimensionality of this set of eight correlated pollutants into a smaller set of uncorrelated variables (principal components (PCs)). Finally, the initial analyses were applied to the PCs to detect underlying patterns of emission sources.

Results: The first analysis detected no association between MD and exposure to any of the pollutants. The second analysis showed non-statistically significant increased risks (OR_{Q4}; IC95%) of high MD were detected in women with higher exposure to SO₂ (1.50; 0.90–2.48), and PM_{2.5} (1.27; 0.77–2.10). In contrast, non-significant ORs < 1 were found in all exposure quartiles for NO (OR_{Q2} = 0.72, OR_{Q3} = 0.68, OR_{Q4} = 0.78), and PM₁₀ (OR_{Q2} = 0.69, OR_{Q3} = 0.82, OR_{Q4} = 0.72). PCA identified two PCs (PC1: “traffic pollution” and PC2: “natural pollution”), and no association was detected between MD and proximity to these two PCs.

Conclusions: In general, our results show a lack of association between residential exposure to specific air pollutants and MD in premenopausal females. Future research is needed to confirm or refute these findings.

Keywords: Breast density, Air pollution, Kriging, Correlation, Principal component analysis, Breast cancer, DDM-Madrid, Long-term exposure

1. Background

Breast cancer is the leading cause of female cancer worldwide [1] and represents a major health problem, as its incidence is expected to rise in the next decade in our country [2]. Regarding environmental exposures, ambient air pollution was classified as carcinogenic to humans [3] because it contains various carcinogens and endocrine-disrupting chemicals (EDCs). In urban zones, traffic-

related air pollution (TRAP) is a significant contributor, which includes substances such as nitrogen oxides (NO_x) and particulate matter (PM) [4].

Some authors have attempted to establish a link between ambient pollutants in urban areas and increased female gynecological cancers [5], and more specifically with breast cancer risk, although with inconclusive findings [6–8]. Nevertheless, some studies have found that women living in environments with high levels of air pollution have an

increased risk of developing breast cancer. This increased risk has been associated with exposure to specific pollutants, such as nitrogen dioxide (NO₂) [9, 10], PM [11], and sulfur dioxide (SO₂) [12]. In the case of SO₂, it has been proposed that its ability to absorb ultraviolet light could interfere with vitamin D synthesis in the skin, leading to vitamin D deficiency and, consequently, an increased risk of breast cancer [13]. In addition, exposure to air pollutants has been shown to be associated with an increased frequency of genetic damage, including changes in DNA methylation patterns and telomere shortening [14–17]. These alterations may influence the expression of genes related to DNA repair, cell cycle control, inflammation and the response to oxidative stress [18]. In particular, exposure to PM, which is composed of a variety of substances with endocrine and carcinogenic properties [19], has been linked to accelerated biological age, as measured by DNA methylation [20].

Mammographic density (MD), the proportion of radiologically dense fibro-glandular tissue in the breast [21], is a well-established risk factor for breast cancer. An important characteristic of this biomarker is its dynamic and modifiable nature, which can be altered by various factors, including parity, body mass index (BMI), age, and menopausal transition [22]. The importance of this biomarker for breast cancer risk is such that it has been established that the risk of breast cancer can be up to 6 times higher in women with MD above 75% compared to those with lower densities (MD < 5%) [21, 23]. Regarding MD and air pollution exposure, the available literature focusing on specific pollutants (PM, NO_x, or ozone (O₃)) is scarce, and shows inconsistent findings [24–30]. Investigating the possible effect of certain air pollutants on MD may help to clarify the causal mechanism of the associations detected with breast cancer found in previous studies. It is also the risk factor with the highest attributable fraction [31], and its inclusion in the breast cancer risk prediction models improves the estimate [32]. Therefore, we aimed to assess whether there is a relationship between MD and exposure to eight air pollutants (SO₂, carbon monoxide (CO), nitrogen monoxide (NO), NO₂, NO_x, PM < 2.5 μm (PM_{2.5}), PM < 10 μm (PM₁₀), and O₃) in premenopausal Spanish women.

2. Methods

2.1. Study population

DDM-Madrid is a Spanish cross-sectional study carried out from 2013 to 2015 [33]. A total of 1466 premenopausal women (39–50 years) were recruited from the Madrid Medical Diagnostic Center during their routine gynecological examinations. Participants were contacted by phone, and 88% of them agreed to participate in the study providing their written informed consents. Trained personnel administered an epidemiological questionnaire on the same day as participants' medical examinations. The comprehensive data collection covered sociodemo-

graphic information, medical history, lifestyle, and residential history. Geographical data were obtained by geocoding each participant's address into EPSG:23030 coordinates. We excluded 24 participants with missing data on the address variable, and restricted the analysis to the 896 women residing in the municipality of Madrid. Madrid is the second most populated city in the EU (3.3 million inhabitants), and it experiences intense road traffic, with high NO₂ levels exceeding EU limit values [34]. After excluding 11 women whose MD could not be measured, 18 participants with analogical images, and 98 participants with missing data in covariates, the final sample comprised 769 women, whose geographic distribution is depicted in Fig. 1.

2.2. Mammographic density assessment

The 2D digital mammograms of both breasts were collected in craniocaudal and mediolateral oblique views. An experienced radiologist assessed the % of MD from the craniocaudal view of the left breast using DM-Scan, a semi-automated computer tool with high validity and reproducibility [35]. Additionally, we employed the American College of Radiology's Breast Imaging-Reporting and Data System (BI-RADS) classification [36], to categorize breast density based on the amount of fibroglandular tissue present. The classification includes: category 1) almost entirely fat (<25% of fibroglandular tissue); category 2) scattered density (25–50%); category 3) heterogeneously dense (51–75%); and category 4) extremely dense (>75%). Finally, we categorized MD into two groups: "low" breast density (categories 1+2) vs. "high" breast density (categories 3+4).

2.3. Air pollution exposure assessment

2.3.1. Air quality data gathering

The Madrid City Council has an Air Quality Monitoring System that comprises 24 automatic remote monitoring stations (12 urban background, 9 traffic, and 3 suburban stations) located throughout all districts. The urban background stations are representative of the exposure of the urban population in general. The traffic stations are located in areas where the pollution level is mainly influenced by emissions from a nearby street or road. The suburban stations are located on the outskirts of the city, where the highest ozone levels are found. Information on the monitoring stations is publicly available on the Madrid City Council's Open Data website [37]. For the present paper, we selected the following eight pollutants, according to similar previous studies [25, 26]: SO₂ (measured by ultraviolet fluorescence at 10 monitoring stations), CO (measured by infrared absorption at 9 stations), NO, NO₂, and NO_x (measured by chemiluminescence at 24 stations), PM_{2.5} (measured by microbalance at 6 stations), PM₁₀ (measured by microbalance at 12 stations); and O₃ (measured by ultraviolet absorption at 14 stations). Hourly data (concentrations) of these pollutants were collected over a 3-year period (2013–2015), and annual averages were cal-

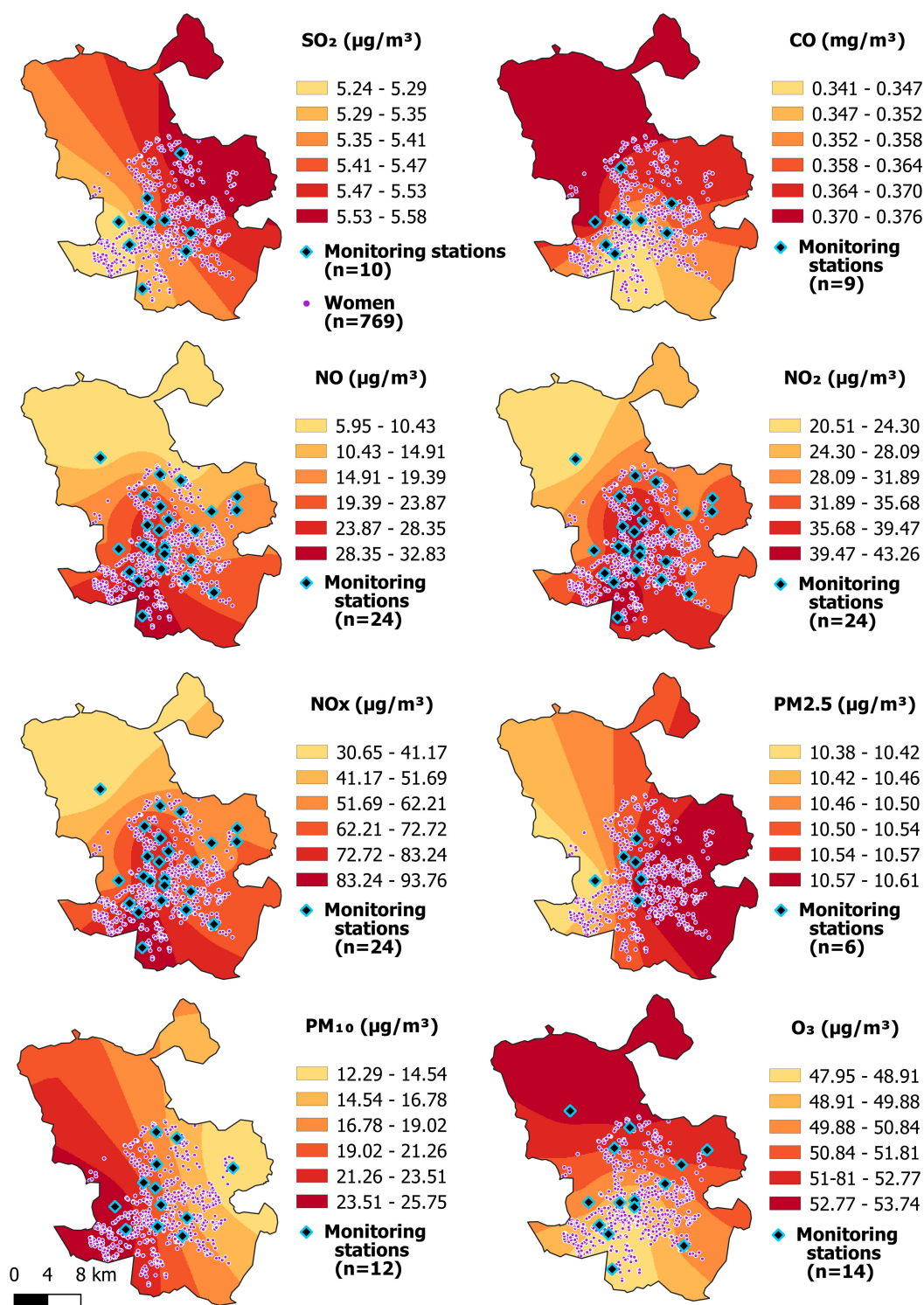


Fig. 1 Concentrations of pollutants interpolated by kriging, and locations of women's addresses and monitoring stations.

culated. It should be noted that data not validated by the City Council were discarded for the calculation of annual averages, as well as those hourly data for which no information was available.

2.3.2 Individual long-term exposure assessment

In order to ascertain the long-term residential exposure of

each woman to the aforementioned pollutants, the monitoring stations were geocoded into EPSG:23030 coordinates, and ordinary kriging was applied to the 3-year annual mean concentrations of each pollutant to interpolate the surface of the municipality of Madrid. Then, each woman was assigned the interpolated concentration for each pollutant corresponding to her residence to estimate

the 3-year annual mean of exposure. Finally, this exposure variable for each pollutant was categorized into quartiles based on its distribution in the study population.

Kriging is a geostatistical method of spatial interpolation of data used to estimate the value of a variable over a continuous spatial field. Ordinary kriging is one of the most widely used kriging methods and uses weighted averages of neighboring points to determine non-measured values or interpolation points. It is considered to be the best linear unbiased estimator [38–40]. Figure 1 shows the interpolated values (contour maps) for each air pollutant in the municipality of Madrid using ordinary kriging, as well as the monitoring stations.

Because air pollutants are often correlated, a correlation analysis was performed using Pearson's correlation coefficients. The Pearson's correlation coefficient, which varies between -1 and $+1$, is a statistical tool commonly used in linear regression that measures the degree of linear association between two variables, and has been used to measure the strength of the relationships among the eight pollutants analyzed in our study. Subsequently, we performed a principal component analysis (PCA), a technique of multivariate data analysis, to reduce the dimensionality of this set of eight correlated exposure variables (interpolated values for each pollutant using ordinary kriging) into a smaller set of new uncorrelated or independent variables called principal components (PCs), with a minimal loss of information [41]. These PCs are linear combinations of the original variables (air pollutants) that allow the visualization of correlations and patterns among the pollutants, thus identifying potential emission sources. PCs with eigenvalues >1 were considered in our analyses (Kaiser or latent root criterion).

2.4. Statistical analyses

We examined descriptive characteristics of the females by calculating absolute figures and percentages. Percent MD was calculated according to these characteristics, and we presented means along with their corresponding standard deviations and 95% confidence intervals (95% CIs). We used two-sided χ^2 test to compare descriptive characteristics between participants with “low” and “high” breast density.

Two analyses were performed to evaluate the association between MD and residential exposure to the selected air pollutants:

- 1) Analysis 1 (MD as a continuous variable). The percentage of MD was log-transformed to achieve the statistical assumptions for linear regression (linearity, homoscedasticity, independence of errors, and normally distributed residuals) [42, 43]. Subsequently, the estimated β coefficients and standard errors were then exponentiated to calculate the relative change in the adjusted geometric mean of the percent MD (e^β), comparing participants among the different exposure categories [44]. Multiple linear regression models were used to estimate e^β and its 95%CI for each quar-

tile of exposure (considering Q1 as the reference), for each pollutant (8 independent models in total). Finally, regression diagnostics were checked for all the models by means of graphical analysis of residuals (standardized residuals vs. fitted values; means and standard deviations of the standardized residuals by deciles of fitted values; and normal Q-Q plots).

- 2) Analysis 2 (MD as a dichotomous variable: “high” breast density/“low” breast density). Multiple logistic regression models were used to estimate odds ratios (ORs) and 95% CIs, which represent the risk of having a higher MD in each quartile of exposure compared to Q1 as the reference quartile, for each pollutant (8 independent models, in total). The dependent variable was categorized as “low” MD = 0 and “high” MD = 1. The main assumptions used in logistic regression are: the conditional distribution of the outcome variable follows a binomial distribution, linearity in the logit (linear relationship between the natural logarithm of the odds and the continuous independent variables included in the model), independence between observations, and the binomial rather than the normal distribution describes the distribution of the errors. Logistic regression is a powerful tool, especially in epidemiologic studies, that allows multiple covariates to be analyzed simultaneously, thus reducing the effect of confounding factors [45]. Finally, the significance of each model used in the analysis was tested by means of the likelihood ratio test, calculating the difference in residuals (deviance) between the model with predictors and the model without predictors (null model).

Finally, the association between MD and residential exposure to PCs (or combined air pollutants) was assessed by applying the methodology used in the two previous analyses (for specific pollutants) to the PCs with eigenvalues >1 obtained from the PCA.

All statistical models were adjusted for recognized and potential confounding factors related to MD [22]: age and BMI (as continuous); and education, alcohol and tobacco consumption, previous breast biopsies, parity, family history of breast cancer, energy intake, and oral contraceptives use (as categorical).

For power calculations, assuming that 50% of women are exposed to high levels of air pollution, and with a sample size of 385 women for each pair of extreme quartiles between which pollution exposure is to be compared, we expected to detect a level of difference in the log-transformed percent MD in each exposure quartile of 0.09, 0.11, and 0.14 with a power of 40, 60%, and 80%, respectively. On the other hand, with a sample size of 300 participants with “low” MD and 100 with “high” MD for each pair of extreme quartiles between which pollution exposure is to be compared, we expected to detect ORs of having “high” breast density in each exposure quartile of 1.49 (or, inversely, 0.67), 1.68 (or, inversely, 0.59), and 1.94 (or, inversely, 0.52) with a power of 40, 60, and 80%, respectively.

Analyses were performed using R 4.3.0 and Stata/IC 16.1 (StataCorp LLC) software, and QGIS 3.28.16.

3. Results

3.1. Main characteristics of the study population (Table 1)

The majority of participants were younger than 45 years (52.1%), had a healthy weight (68.1%), had a university grade (64.1%), consumed <10 g/day of alcohol (67.0%), and had no previous biopsies (90.5%). MD (mean ± SD) was higher in women younger than 45 years old (36.8 ± 17.6), with a normal BMI (39.3 ± 16.8), having a university education (36.1 ± 17.9), never smokers (37.2 ± 18.6), and nulliparous (37.7 ± 18.5). Finally, the group with “low” density had a statistically significant higher number of women with obesity or overweight (BMI ≥ 25) and no previous biopsies (*p*-values < 0.001 and 0.034, respectively). Conversely, the number of nulliparous participants and the number of women who had never used oral contraceptives was higher in the group with “high” density (*p*-values = 0.005 and 0.006, respectively).

3.2. Results of the analysis 1 (MD as a continuous variable) (Table 2)

No association was detected between MD and exposure to any of the pollutants, where the results were very close to the unity for all exposure quartiles. In the case of NO, NO_x, and PM₁₀, the results showed non-statistically significant decreased MD in all exposure quartiles, with no dose-response trends (*p*-trends = 0.646, 0.798, and 0.470, respectively).

3.3. Results of the analysis 2 (MD as a dichotomous variable: “high” density/“low” density) (Table 3)

Our results showed no association between risk of higher MD and exposure to any of the pollutants. However, non-statistically significant ORs > 1 were found in women exposed to all exposure quartiles for SO₂, especially in the highest exposure quartile (OR_{Q2} = 1.32, OR_{Q3} = 1.10, OR_{Q4} = 1.50), and CO (OR_{Q2} = 1.14, OR_{Q3} = 1.04, OR_{Q4} = 1.05). On the other hand, O₃ and PM_{2.5} showed ORs > 1 in the highest exposure quartile (OR_{Q4} = 1.10 and 1.27, respectively), and ORs ≤ 1 in the remaining exposure quartiles. Finally, non-statistically significant ORs < 1 were detected in all exposure quartiles for NO, NO₂, NO_x, and PM₁₀.

3.4. Results of the correlation and principal component analyses

Correlation analysis (Fig. 2) showed positive Pearson’s correlation coefficients >0.80 between SO₂ and PM_{2.5} (0.84), CO and O₃ (0.94), NO and NO_x (0.95), and NO₂ and NO_x (0.89). On the contrary, negative Pearson’s correlation coefficients lower than −0.80 were detected between SO₂ and PM₁₀ (−0.95), NO and O₃ (−0.85), and PM_{2.5} and PM₁₀ (−0.89).

Table 1 Descriptive characteristics of the participants.

Characteristics	All participants		Participants with “low” breast density		Participants with “high” breast density		<i>p</i> -value ^b
	n (%)	Mammographic density (%) Mean (95%CI) SD ^a	n (%)	Mammographic density (%) Mean (95%CI) SD ^a	n (%)	Mammographic density (%) Mean (95%CI) SD ^a	
Total	769 (100.0)	35.0 (33.8; 36.2) 17.4	597 (100.0)	27.9 (26.9; 28.9) 11.9	172 (100.0)	59.6 (58.3; 61.0) 8.8	
Age							
<45	401 (52.1)	36.8 (35.1; 38.5) 17.6	300 (50.3)	28.9 (27.6; 30.2) 11.6	101 (58.7)	60.3 (58.4; 62.2) 9.8	0.061
≥45	368 (47.9)	33.0 (31.3; 34.8) 17.0	297 (49.7)	26.9 (25.5; 28.3) 12.2	71 (41.3)	58.8 (57.1; 60.4) 7.3	
Body mass index (kg/m²)							
<18.5	15 (2.0)	38.4 (32.0; 44.8) 12.6	12 (2.0)	34.3 (28.4; 40.1) 10.3	3 (1.7)	55.1 (51.0; 59.2) 3.6	<0.001
18.5–24.9	524 (68.1)	39.3 (37.9; 40.8) 16.8	377 (63.1)	31.1 (30.0; 32.2) 10.9	147 (85.5)	60.4 (59.0; 61.9) 9.0	
25–29.9	161 (20.9)	27.0 (24.8; 29.3) 14.7	143 (24.0)	23.4 (21.6; 25.2) 11.0	18 (10.5)	55.5 (52.4; 58.7) 6.8	
≥30	69 (9.0)	20.0 (16.7; 23.3) 14.0	65 (10.9)	18.0 (15.2; 20.9) 11.8	4 (2.3)	52.3 (50.4; 54.3) 2.0	
Education							
Primary school or less	26 (3.4)	30.8 (23.7; 38.0) 18.7	20 (3.3)	23.7 (17.2; 30.2) 14.9	6 (3.5)	54.7 (51.2; 58.3) 4.4	0.441
Secondary school	250 (32.5)	33.3 (31.3; 35.3) 16.2	201 (33.7)	27.4 (25.8; 29.0) 11.6	49 (28.5)	57.4 (55.3; 59.5) 7.6	
University grade	493 (64.1)	36.1 (34.5; 37.7) 17.9	376 (63.0)	28.4 (27.2; 29.6) 11.9	117 (68.0)	60.8 (59.2; 62.5) 9.2	

Table 1 (Continued.)

Characteristics	All participants			Participants with “low” breast density			Participants with “high” breast density			p-value ^b
	n (%)	Mean (95%CI)	SD ^a	n (%)	Mean (95%CI)	SD ^a	n (%)	Mean (95%CI)	SD ^a	
Alcohol consumption (g/day)										
Never	134 (17.4)	34.5 (31.5; 37.4)	17.4	104 (17.4)	27.4 (25.1; 29.7)	11.9	30 (17.4)	59.1 (56.1; 62.1)	8.3	0.647
<10	515 (67.0)	35.2 (33.7; 36.7)	17.7	396 (66.3)	27.8 (26.6; 29.0)	11.9	119 (69.2)	59.8 (58.1; 61.5)	9.2	
≥10	120 (15.6)	34.8 (31.9; 37.8)	16.5	97 (16.3)	29.0 (26.6; 31.4)	12.1	23 (13.4)	59.5 (56.4; 62.5)	7.5	0.098
Tobacco consumption										
Never	297 (38.6)	37.2 (35.1; 39.3)	18.6	220 (36.9)	28.3 (26.8; 29.9)	11.4	77 (44.8)	62.6 (60.3; 64.8)	10.0	
Former smoker	262 (34.1)	33.4 (31.4; 35.3)	16.1	214 (35.8)	28.1 (26.4; 29.8)	12.4	48 (27.9)	56.8 (54.8; 58.9)	7.4	
Current smoker	210 (27.3)	33.9 (31.6; 36.2)	17.0	163 (27.3)	27.1 (25.2; 28.9)	12.2	47 (27.3)	57.7 (55.9; 59.6)	6.5	0.034
Previous breast biopsies										
Yes	73 (9.5)	40.4 (36.5; 44.3)	16.9	49 (8.2)	31.1 (27.9; 34.4)	11.7	24 (14.0)	59.4 (56.4; 62.3)	7.3	
No	696 (90.5)	34.4 (33.1; 35.7)	17.4	548 (91.8)	27.6 (26.6; 28.6)	11.9	148 (86.0)	59.7 (58.2; 61.1)	9.1	0.005
Parity										
Nulliparous	225 (29.3)	37.7 (35.3; 40.1)	18.5	157 (26.3)	28.1 (26.2; 30.0)	12.3	68 (39.5)	59.8 (57.8; 61.8)	8.5	
1	171 (22.2)	34.4 (31.7; 37.1)	18.0	134 (22.5)	27.1 (25.1; 29.0)	11.4	37 (21.5)	61.1 (57.7; 64.6)	10.8	
2	335 (43.6)	33.8 (32.0; 35.5)	16.4	273 (45.7)	28.2 (26.7; 29.6)	12.1	62 (36.1)	58.4 (56.5; 60.4)	8.0	
>2	38 (4.9)	32.6 (27.6; 37.5)	15.6	33 (5.5)	28.3 (24.4; 32.3)	11.6	5 (2.9)	60.7 (54.6; 66.8)	7.0	0.451
Family history of breast cancer										
None	589 (76.6)	35.2 (33.8; 36.6)	17.3	455 (76.2)	28.1 (27.0; 29.2)	12.0	134 (77.9)	59.3 (57.9; 60.8)	8.6	
Second degree only	128 (16.6)	34.7 (31.5; 37.9)	18.4	98 (16.4)	26.8 (24.4; 29.1)	12.0	30 (17.4)	60.6 (57.1; 64.1)	9.7	
First degree	52 (6.8)	33.4 (28.9; 38.0)	16.7	44 (7.4)	28.3 (24.8; 31.8)	11.8	8 (4.7)	61.6 (54.7; 68.5)	10.0	0.333
Energy intake (kcal/day)^c										
>2172.5	255 (33.1)	34.9 (32.7; 37.1)	18.1	192 (32.2)	26.8 (25.1; 28.5)	12.1	63 (36.6)	59.6 (57.7; 61.6)	7.9	
1657.6–2172.5	258 (33.6)	35.9 (33.8; 37.9)	16.5	208 (34.8)	30.0 (28.4; 31.6)	11.6	50 (29.1)	60.3 (57.3; 63.3)	10.7	
<1657.6	256 (33.3)	34.2 (32.0; 36.4)	17.7	197 (33.0)	26.8 (25.1; 28.4)	12.0	59 (34.3)	59.1 (57.0; 61.1)	8.1	0.006
Oral contraceptives consumption										
Current use	26 (3.4)	32.3 (25.5; 39.2)	17.9	23 (3.8)	27.4 (23.1; 31.7)	10.5	3 (1.7)	70.0 (47.5; 92.4)	19.8	
Past use	423 (55.0)	33.5 (31.9; 35.1)	16.5	343 (57.5)	27.7 (26.5; 29.0)	12.2	80 (46.5)	58.1 (56.5; 59.7)	7.2	
Never	320 (41.6)	37.2 (35.2; 39.2)	18.4	231 (38.7)	28.2 (26.7; 29.7)	11.8	89 (51.8)	60.7 (58.7; 62.6)	9.4	

^a Standard deviation.

^b Two-sided χ^2 test was used to compare descriptive characteristics between participants with “low” and “high” breast density.

^c Variable in tertiles.

Table 2 Association between log-transformed percentage of mammographic density and residential exposure to specific air pollutants.

Pollutant exposure quartile	n	e ^β (95%CI) ^a	p-trend
SO₂ (μg/m³)			
Q1 (reference)	194	1	
Q2	189	1.00 (0.89–1.13)	
Q3	190	0.96 (0.85–1.08)	
Q4	196	1.05 (0.93–1.18)	0.571
CO (mg/m³)			
Q1 (reference)	189	1	
Q2	192	0.98 (0.87–1.10)	
Q3	192	1.01 (0.89–1.14)	
Q4	196	1.01 (0.89–1.14)	0.775
NO (μg/m³)			
Q1 (reference)	198	1	
Q2	185	0.90 (0.80–1.02)	
Q3	197	0.93 (0.82–1.04)	
Q4	189	0.96 (0.86–1.09)	0.646
NO₂ (μg/m³)			
Q1 (reference)	197	1	
Q2	193	0.97 (0.86–1.09)	
Q3	190	1.03 (0.91–1.16)	
Q4	189	0.96 (0.85–1.08)	0.761
NO_x (μg/m³)			
Q1 (reference)	195	1	
Q2	189	0.93 (0.83–1.05)	
Q3	196	0.96 (0.85–1.07)	
Q4	189	0.98 (0.87–1.10)	0.798
PM_{2.5} (μg/m³)			
Q1 (reference)	197	1	
Q2	188	0.99 (0.88–1.11)	
Q3	188	0.98 (0.87–1.10)	
Q4	196	1.02 (0.91–1.15)	0.803
PM₁₀ (μg/m³)			
Q1 (reference)	192	1	
Q2	195	0.97 (0.86–1.09)	
Q3	189	0.94 (0.83–1.06)	
Q4	193	0.96 (0.85–1.08)	0.470
O₃ (μg/m³)			
Q1 (reference)	188	1	
Q2	198	1.04 (0.92–1.17)	
Q3	185	1.00 (0.88–1.13)	
Q4	198	1.01 (0.90–1.14)	1.00

^a e^β estimated using multiple linear regression models, adjusted for age, body mass index, education, alcohol consumption, tobacco consumption, previous breast biopsies, parity, family history of breast cancer, energy intake, and oral contraceptives consumption (an independent model for each pollutant).

Table 3 ORs of having “high” breast density and residential exposure to specific air pollutants.

Pollutant exposure quartile	“Low” breast density (n)	“High” breast density (n)	OR (95%CI) ^a	p-trend
SO₂ (μg/m³)				
Q1 (reference)	155	39	1	
Q2	147	42	1.32 (0.78–2.25)	
Q3	153	37	1.10 (0.64–1.89)	
Q4	142	54	1.50 (0.90–2.48)	0.197
CO (mg/m³)				
Q1 (reference)	150	39	1	
Q2	149	43	1.14 (0.67–1.92)	
Q3	147	45	1.04 (0.62–1.76)	
Q4	151	45	1.05 (0.62–1.77)	0.951
NO (μg/m³)				
Q1 (reference)	145	53	1	
Q2	148	37	0.72 (0.43–1.20)	
Q3	155	42	0.68 (0.41–1.13)	
Q4	149	40	0.78 (0.47–1.30)	0.296
NO₂ (μg/m³)				
Q1 (reference)	150	47	1	
Q2	149	44	0.87 (0.52–1.44)	
Q3	145	45	0.98 (0.59–1.62)	
Q4	153	36	0.71 (0.42–1.21)	0.298
NO_x (μg/m³)				
Q1 (reference)	146	49	1	
Q2	149	40	0.77 (0.46–1.29)	
Q3	154	42	0.72 (0.43–1.20)	
Q4	148	41	0.83 (0.50–1.39)	0.432
PM_{2.5} (μg/m³)				
Q1 (reference)	153	44	1	
Q2	146	42	1.00 (0.60–1.67)	
Q3	152	36	0.86 (0.51–1.46)	
Q4	146	50	1.27 (0.77–2.10)	0.455
PM₁₀ (μg/m³)				
Q1 (reference)	142	50	1	
Q2	157	38	0.69 (0.41–1.15)	
Q3	146	43	0.82 (0.49–1.36)	
Q4	152	41	0.72 (0.43–1.19)	0.303
O₃ (μg/m³)				
Q1 (reference)	147	41	1	
Q2	155	43	0.88 (0.52–1.48)	
Q3	146	39	0.80 (0.47–1.36)	
Q4	149	49	1.10 (0.66–1.83)	0.774

^a ORs estimated using multiple logistic regression models, adjusted for age, body mass index, education, alcohol consumption, tobacco consumption, previous breast biopsies, parity, family history of breast cancer, energy intake, and oral contraceptives consumption (an independent model for each pollutant).

The PCA identified two PCs with eigenvalues >1 (Table 4), accounting for 92.19% of the total variance (69.35% corresponding to PC1 and 22.84% corresponding to PC2). The PCA loadings, which indicate the relative weight of each pollutant within the corresponding PC, are presented in Table 5: the first PC (PC1) shows positive loadings for NO, NO₂, NO_x, and PM₁₀, and negative loadings for the remaining pollutants. This PC would be associated with pollutants related to traffic pollution, inasmuch as it would represent the origin of nitrogen oxides and their subsequent photochemical reaction to produce ozone from

road traffic [39]. The second PC (PC2) shows positive loadings for PM₁₀, CO, and O₃, and negative loadings for the remaining pollutants. In particular, this PC discriminates between PM₁₀ (loading = 0.3874) and PM_{2.5} (loading = -0.6135); this fact could suggest the influence of other sources of PM other than photochemical reactions, such as episodes of Saharan dust intrusions and/or soil abrasion [39, 46, 47]. On the other hand, one of the main sources of CO is from wildfires [48, 49] and, in this regard, during the period 2013–2015 there were 986 wildfires in the Madrid region, affecting an area of 2285 ha [50].

Finally, O₃ is a pollutant that is not directly released by primary sources, but is formed through complex reactions in the atmosphere driven by the energy transferred to NO₂ molecules when they absorb light from solar radiation [49]. Therefore, PC2 would be associated with “natural pollution”, i.e., pollution of natural origin that is not due to anthropogenic sources resulting from human activities.

Table 6 shows the association between the log-transformed percentage of MD and proximity to PC1 (“traffic pollution”) and PC2 (“natural pollution”). No association was detected with the two PCs, where the results were very close to the unity in all exposure quartiles of PC1. In the

case of PC2, the results showed non-statistically significant decreased MD in exposure quartiles Q2 and Q3 ($e^{\beta} = 0.89$ in both cases).

Finally, our results showed no association between risk of having “high” MD and exposure to any of the PCs (Table 7). Non-statistically significant ORs < 1 were detected in all exposure quartiles for PC1, and in Q2 and Q3 for PC2.

4. Discussion

This study examines the association between MD and exposure to eight specific ambient pollutants in premenopausal women residing in the municipality of Madrid. MD, one of the most important biomarkers of breast cancer risk, is characterized by its modifiable nature, which makes it susceptible to the influence of different exposures. This vulnerability led us to examine the possible effect of exposure to air pollutants. In general, our results revealed no association between residential exposure to air pollutants and MD in premenopausal participants. However, we detected non-statistically significant increased risks of having high MD in women exposed to the highest concentrations of SO₂, PM_{2.5}, and O₃, and conversely, non-statistically significant decreased risks of having high MD in women exposed to NO, NO₂, NO_x, and PM₁₀. On the other hand, PCA did not detect an association between MD and residential exposure to “traffic” and “natural” pollution.

Breast tissue changes may be influenced by several factors. Research has linked ambient pollution to an increased risk of breast tumors in urban environments [8]. On the other hand, some studies have also shown that females residing in urban zones have higher MD compared to those in rural zones [51, 52]. Furthermore, MD has been associated with ambient pollution [29, 30]. TRAP present in urban areas includes substances such as EDCs and carcinogens, which have the potential to influence the endocrine system and potentially may impact mammary gland development, ultimately leading to possible modifications in MD [53]. Lastly, it is important to mention that MD, by its dynamic nature, is influenced by relatively recent environmental exposures, meaning that MD at the time of assessment is likely to reflect exposure to pollutants during that specific time period. Therefore, it is critical to select air pollutant exposure data from the period closest to the time of data collection.

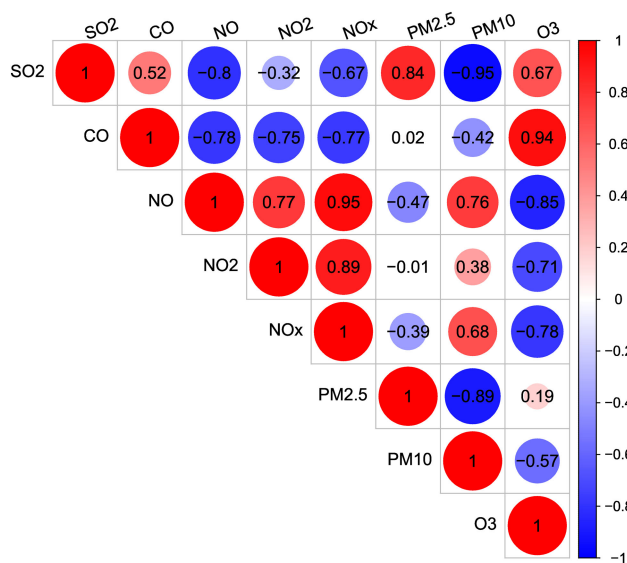


Fig. 2 Correlation matrix among air pollutants.

Table 4 Eigenvalues, explained variance, and cumulative percentage of the principal components (PCs).

PC number	Eigenvalue	Explained variance (%)	Cumulative percentage
1	5.548	69.35	69.35
2	1.827	22.84	92.19
3	0.462	5.77	97.96
4	0.083	1.04	99.00
5	0.054	0.68	99.68
6	0.013	0.17	99.85
7	0.010	0.12	99.97
8	0.002	0.03	100.00

Table 5 PCA loadings.

Pollutant	PC1	PC2	PC3	PC4	PC5	PC6	PC7	PC8
SO ₂	-0.3642	-0.3492	0.2650	-0.1618	-0.0666	0.5133	-0.5131	0.3433
CO	-0.3438	0.3542	0.4333	0.1913	0.6757	0.1368	0.0616	-0.2301
NO	0.4139	-0.0489	0.0988	0.6592	0.1977	-0.2016	-0.3771	0.4000
NO ₂	0.3177	-0.3901	0.5571	-0.4090	0.2235	-0.2556	0.3220	0.2243
NO _x	0.3973	-0.1377	0.4041	0.2533	-0.3324	0.4058	0.0109	-0.5661
PM _{2.5}	-0.2311	-0.6135	-0.1453	0.0756	0.2125	-0.4044	-0.2860	-0.5030
PM ₁₀	0.3550	0.3874	0.0421	-0.5038	0.1056	-0.1364	-0.6272	-0.2125
O ₃	-0.3737	0.2295	0.4856	0.1082	-0.5322	-0.5148	-0.1079	0.0183

Table 6 Association between log-transformed percentage of mammographic density and residential exposure to principal components (PCs).

PC exposure quartile	n	e ^β (95%CI) ^a	p-trend
PC1: “traffic pollution”			
Q1 (reference)	198	1	
Q2	187	0.94 (0.83–1.06)	
Q3	193	0.97 (0.86–1.09)	
Q4	191	1.01 (0.89–1.13)	0.828
PC2: “natural pollution”			
Q1 (reference)	189	1	
Q2	190	0.89 (0.79–1.00)	
Q3	194	0.89 (0.79–1.00)	
Q4	196	0.99 (0.88–1.11)	0.850

^a e^β estimated using multiple linear regression models, adjusted for age, body mass index, education, alcohol consumption, tobacco consumption, previous breast biopsies, parity, family history of breast cancer, energy intake, and oral contraceptives consumption (an independent model for each PC).

Table 7 ORs of having “high” breast density and residential exposure to principal components (PCs).

PC exposure quartile	“Low” breast density (n)	“High” breast density (n)	OR (95%CI) ^a	p-trend
PC1: “traffic pollution”				
Q1 (reference)	149	49	1	
Q2	147	40	0.85 (0.51–1.42)	
Q3	152	41	0.81 (0.49–1.35)	
Q4	149	42	0.88 (0.53–1.46)	0.589
PC2: “natural pollution”				
Q1 (reference)	148	41	1	
Q2	149	41	0.91 (0.54–1.53)	
Q3	151	43	0.98 (0.58–1.63)	
Q4	149	47	1.03 (0.62–1.73)	0.835

^a ORs estimated using multiple logistic regression models, adjusted for age, body mass index, education, alcohol consumption, tobacco consumption, previous breast biopsies, parity, family history of breast cancer, energy intake, and oral contraceptives consumption (an independent model for each PC).

Although air quality data gathering was collected ten years ago, air pollution in urban areas is an ongoing problem. With respect to the temporal evolution in the air pollutant levels in Madrid city during the period 2013–2023, some pollutants have shown decreasing average changes in their annual average levels, such as PM_{2.5} (−6.1% between 2013 and 2023), and NO₂ (−21.5%) [37], probably due to the traffic restrictions imposed in Madrid in recent years and the mobility restrictions during the COVID-19 pandemic [54, 55]. However, other pollutants have shown increased average changes in their annual levels over the same period, such as CO (1.3%), PM₁₀ (8.8%), and O₃ (12.4%) [37].

This is the first study conducted in Spain that analyzes the potential association between MD and residential exposure to concentrations of specific air pollutants. Our findings in premenopausal women support those obtained in the work by Eslami et al. [25], carried out in Iranian

women, the only cross-sectional epidemiological study published to date on MD and exposure to concentrations of specific air pollutants, such as ours. In that study, conducted in women attending two hospitals in Tehran for screening mammography, the total number of eligible participants was 791, although only 396 were premenopausal. In contrast, our study had a significantly higher number of premenopausal women (n = 769). In terms of air quality data gathering, the Iranian study included 17 monitoring stations and six ambient air pollutants, whereas our study included more monitoring stations (n = 24) and air pollutants (n = 8). However, it is important to note that not all air pollutants were measured at each monitoring station, which is a limitation of our study. Both studies used a 3-year annual mean of exposure for each pollutant to assess individual long-term exposure. With respect to the statistical analyses, the Iranian study used multiple logistic regression models to evaluate the potential relationship between MD and exposure to the selected pollutants. The dependent variable MD was categorized into “high” density vs. “low” density (based on the BI-RADS classification), and the exposure variable for each pollutant was considered as continuous. In our study, we have attempted to replicate the methodology used by Eslami et al. with the same MD classification (“high” DM vs. “low” DM). However, we have refined the exposure measurement by categorizing the exposure variable for each pollutant into quartiles (analysis 2). In addition, it is important to note that we performed an additional analysis (analysis 1) using multiple linear regression models, which were not included in the work by Eslami et al., considering the log-transformed MD as a continuous variable. This analysis offers a more detailed explanation of the association between MD and exposure to ambient pollution. Finally, the previous two analyses were applied to the PCs (linear combinations of the original pollutants), with the purpose of detecting underlying patterns of common emission sources. The subsequent subsections provide a comparison of the findings from both studies and other studies.

4.1. Nitrogen oxides (NO_x, NO₂, and NO)

NO_x, stemming from combustion processes in urban areas, comprises NO and NO₂ (among others), primarily emitted by motorized traffic, making it a reliable marker for vehicle emissions. Specifically, NO₂ is considered a dependable indicator of traffic, since over 75% of atmospheric NO₂ in urban areas is attributed to road traffic [37]. Furthermore, research confirms that the highest levels of NO_x are detected in the vicinity of highways, which are characterized by higher traffic flows [56]. Only two studies have analyzed the relationship between MD and exposure to nitrogen oxides in premenopausal participants. Eslami et al. [25] found a non-statistically significant increased risk of high MD for each unit increase in NO₂ (ppb) (OR = 1.035). In our study, we categorized NO₂ exposure (in µg/m³) into quartiles, and we observed non-statistically significant ORs < 1 in all exposure quartiles. On the other

hand, Huynh et al. [26] analyzed 839 premenopausal women included in the Danish Diet, Cancer and Health cohort (1993–2001) who attended mammographic screening in Copenhagen and, in consonance with our study, found a non-statistically significant inverse association between having mixed/dense MD and exposure to NO_x (OR = 0.92, per $20 \mu\text{g}/\text{m}^3$) and NO_2 (OR = 0.77, per $10 \mu\text{g}/\text{m}^3$). The findings regarding the lack of association between long-term exposure to NO_x/NO_2 and MD in premenopausal women are consistent with those observed in some studies on breast cancer, where the authors have found no association between exposure to NO_2 and breast cancer risk in premenopausal women [57], although inconsistent with others that have reported such an association [7, 58].

4.2. Ozone (O_3)

In urban environments with elevated NO levels, O_3 is rapidly consumed by the oxidation of NO to NO_2 . As a result, O_3 levels tend to be lower in urban traffic zones than in less polluted environments. The latter receive, during the transport of air masses, the O_3 generated in polluted urban and industrial areas, which is not consumed, due to the lack of local NO [59]. Thus, O_3 can be considered as an indicator of air quality that has been hardly evaluated. Eslami et al. [25] did not find any significant relationship between O_3 exposure and high MD (OR = 0.993), in line with that observed in our study. On the other hand, a study conducted in the US between 2001 and 2009 within a large population-based prospective cohort (the Breast Cancer Surveillance Consortium), showed statistically significant increased risks of having a mammogram classified as BI-RADS III (heterogeneously dense breasts) or IV (extremely dense breasts) in premenopausal women exposed to O_3 for all exposure quartiles, except for women classified as BI-RADS IV in the highest quartile, where no association was found [30]. Finally, regarding breast cancer risk, although a previous study found no association with O_3 exposure in premenopausal women [60]. It has been proposed that ozone might exhibit anticancer properties by promoting the generation of reactive oxygen species and relieving hypoxia [61].

4.3. Sulfur dioxide (SO_2)

SO_2 is used to assess ambient air quality, and its primary source is the combustion of fossil fuels in the industrial sector. Eslami et al. [25] found no statistically significant association between SO_2 exposure and MD in premenopausal participants (OR = 0.683). However, our study suggests an increase in MD in premenopausal females exposed to higher levels of SO_2 . Regarding breast cancer, an Iranian study found no association between a higher incidence of breast tumors in premenopausal females and long-term exposure to SO_2 [62].

4.4. Particulate matter (PM_{10} and $\text{PM}_{2.5}$)

PM is comprised of a wide variety of compounds. PM_{10}

and $\text{PM}_{2.5}$ have a greater capability to enter the body through the respiratory tract [63]. With regard to MD, Eslami et al. [25] detected a 23% non-statistically significant increased risk of high MD in premenopausal females exposed to $\text{PM}_{2.5}$, a result similar to that observed in our study (27%). They also detected a non-significant decrease in MD with PM_{10} exposure, again in line with our findings. On the other hand, Yaghjyan et al. [30] detected that premenopausal women exposed to the highest exposure quartiles of $\text{PM}_{2.5}$ had higher odds of their mammograms being classified as BI-RADS III, although not consistently as BI-RADS IV. Finally, a nested case-control study carried out within two cohorts showed no association between residential exposure to PM (including $\text{PM}_{2.5}$ and PM_{10}) and MD [24]. In relation to breast tumors in premenopausal females and PM exposure, the findings are inconclusive [6, 60].

4.5. Limitations and strengths

A limitation of our work is the no monitoring of changes in MD over time, due to the cross-sectional nature of the study. Moreover, the lack of residential data during vulnerable stages of the female participants regarding breast tissue changes, such as menarche or pregnancy, which hinders our ability to identify potential associations during these stages. Additionally, participants were recruited from a single center in the municipality of Madrid, which may limit the external validity of our findings. This also implied that the variation in exposure levels of some pollutants was small, limiting the ability to detect possible associations. In this regard, it should also be noted that the statistical power of the study was rather limited, probably due to the small sample size, so that an association between exposure to air pollutants and MD in premenopausal women cannot be ruled out. However, previous studies on exposure to pollutants such as CO, SO_2 , PM_{10} , NO_2 , or NO_x also point to a null association with MD in premenopausal women [24–26]. On the other hand, exposure assessment error is a challenge in spatial and spatiotemporal modeling. Given the impossibility of collecting individual exposure data, we have used interpolated data on exposure to individual pollutants; these data may be subject to both classical error (in which the measured value is an error-prone version of the true exposure, since random noise is added to a correctly measured exposure) and Berkson error (in which the measured value is a smoothed version of the true exposure, since part of the true exposure variability is not captured by the measurement process). A combination of these two errors could explain the lack of statistical associations [64, 65]. Additionally, although the models were adjusted for potential confounders that could modify MD, other environmental exposures may have affected the detection of potential associations between MD and air pollution exposures in their residential areas. Despite adjustment for a wide variety of potential risk factors, residual unmeasured confounders related to air pollution or MD may have affected our results (or may have interfered with the lack of

significant associations). Lastly, some air pollutants, such as PM_{2.5} and CO, were measured only at a limited number of monitoring stations; therefore, the interpolation of their values by kriging may not be very accurate.

On the contrary, the main strength of our work lies in its novelty, since it is the first work conducted in Spain on residential exposure to concentrations of specific ambient pollutants and MD. On the other hand, the air pollution data were obtained from monitoring stations of the official register of Madrid city. Furthermore, a professional radiologist expert in the interpretation of mammograms, who demonstrated high internal consistency, assessed mammographic density on a continuous scale using DM-Scan, a validated computer-assisted method [66]. Finally, two methodological approaches were used in the analyses, considering MD as both a continuous and a dichotomous variable, and classifying the exposure variable for each air pollutant into quartiles to refine the measurement of exposure to air pollution. In addition, the potential correlations among pollutants were considered, and PCA was applied to detect underlying patterns of emission sources.

5. Conclusions

In general, our results show a lack of association between exposure to specific air pollutants and MD in premenopausal females in the municipality of Madrid. Further research is necessary to refute/confirm these findings, since MD is a modifiable biomarker for breast tumors, and knowing the relationship between them and the environmental factors that could modify it is fundamental to design prevention strategies.

Abbreviations

EDCs: Endocrine-disrupting chemicals; TRAP: Traffic-related air pollution; NO_x: Nitrogen oxides; PM: Particulate matter; NO₂: Nitrogen dioxide; SO₂: Sulfur dioxide; MD: Mammographic density; BMI: Body mass index; O₃: Ozone; CO: carbon monoxide; NO: Nitric oxide/nitrogen monoxide; PM_{2.5}: Particulate matter <2.5 µm; PM₁₀: Particulate matter <10 µm; BI-RADS: Breast Imaging-Reporting and Data System; PCA: Principal component analysis; PCs: Principal components; 95% CIs: 95% confidence intervals; ORs: Odds ratios; ISCIII: Carlos III Institute of Health.

Declarations

Ethics approval and consent to participate

The DDM-Madrid project adhered to the guidelines set forth in the Declaration of Helsinki and received formal approval from the Ethics and Animal Welfare Committee of the Carlos III Institute of Health (No PI: CEI PI 02_2012-v2). All women provided written informed consent prior to the inclusion in this study.

Consent for publication

The data that support the findings are not publicly available due to ethical consideration (privacy of research participants).

Availability of data and material

Not applicable.

Competing interests

The authors declare no competing interests.

Funding

This work was supported by the *Instituto de Salud Carlos III* (ISCIII)/Carlos III Institute of Health (EPY-505/19-PFIS, AESI PI15CIII/00013, and AESI PI15CIII/00029), and the *Sociedad Española de Epidemiología* (SEE)/ Spanish Society of Epidemiology ("Lola Cuevas" Award 2024).

Author's contributions

JG-P and VL conceived, designed, and supervised the study. JG-P and TJ conducted the statistical analyses. TJ wrote the original draft. All authors read, revised, and approved the final manuscript.

Acknowledgments

The work presents independent research, and the views expressed are those of the authors and not necessarily those of the ISCIII.

Received: 10 July 2024, Accepted: 2 October 2024

Published online: 22 November 2024

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